

# **Monetary Policy Rules When UIP Conditions and Policy Trilemma do not Hold: the Case of Nepal\***

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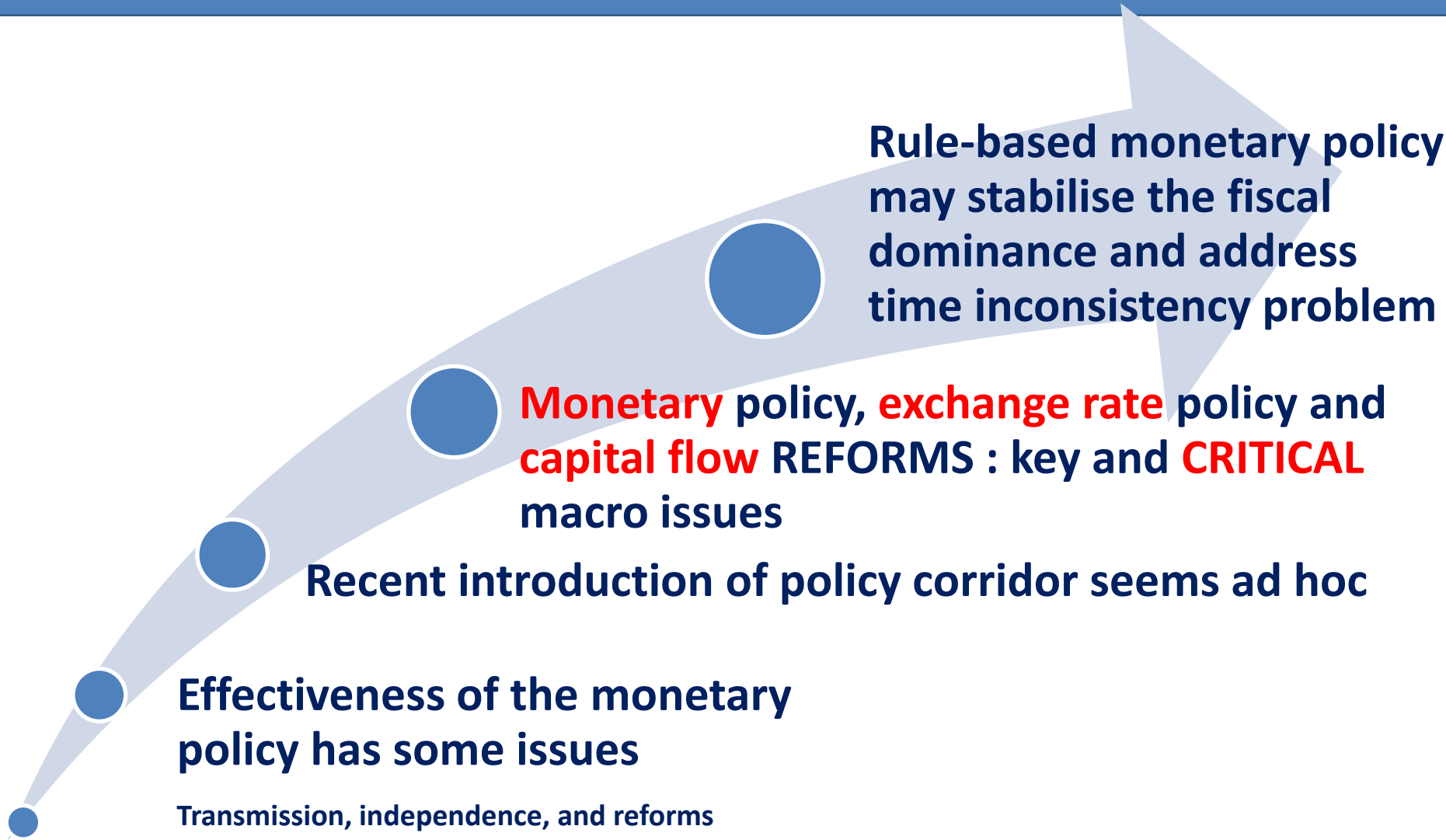
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# Presentation Outline

- **Background**
- **Research objectives and summary gaps**
- **Analytical models**
- **Data**
- **Estimation and Findings**
- **Discussions/Policy Implications**

# Background



Nepal adopts **fixed exchange rate with India** historically

Since 1950s, latest revisit in level of peg: 1993 February

# Summary Research Gaps

- The implication of Taylor (1993) rule has asymmetric implications to economies
- Nepal has a unique economic/political/geographic characteristic that cannot replicate the existing findings
- Our earlier studies (Bhatta et al 2021a, Bhatta et al 2021b) confirm both the UIP and trilemma conditions do not exist in Nepal, that requires a re-designed policy rule

# Research Objectives

- i) explore the factors influencing the short-term interest rate;
- ii) investigate whether the modified Taylor rule holds
- iii) discuss the policy implications

# Theoretical Context : 3 Foundations

## ISLM Model

### Keynesian economic theory

- **Mundell (1963) and Fleming (1962): Impossible Trinity**

## Fisher Effect

### International Fisher Effect:

- Interaction of monetary policy and exchange rate.

**Uncovered Interest parity:**  $e_t^E - e_t = i_t - i_t^*$

## Monetary Policy Rule

### New Keynesian economics

**Taylor rule:**  $i = \pi^* + r^* + \beta (\pi^e - \pi) + \chi (Y - \bar{Y})$

# Modified Taylor Rule

## Uncovered interest parity model

- Backward-looking strategy as  $b_1 \Delta e_t$  and forward-looking strategy as  $(1-b_1)(i_t - i_t^* - \text{prem}_t)$  (Berg et al. 2006a, 2006b)

$$e_{t+1} - e_t = b_1(e_t - e_{t-1}) + (1 - b_1)(i_t - i_t^* - \text{prem}_t) + \varepsilon_t$$

**Monetary policy reaction function (for free market)** Berg et al. 2006a, 2006b

$$i_t = f_1 i_{t-1} + (1 - f_1)(i_t^n + f_2(\pi_t^e - \pi^T) + f_3 \hat{y}_t) + \varepsilon_t^i$$

## Extended reaction function for Nepal :

- $i_t = c_1(\Delta e_{t+1} + i_t^* + \text{prem}_t) + (1 - c_1)(f_1 i_{t-1} + (1 - f_1)(i_t^n + f_2(\pi_{t+1}^e - \bar{\pi}) + f_3 \hat{y}_t)) + \varepsilon_t^i$ 
  - If  $\Delta e_{t+1} = 0$ :  $i_t$  determined by  $i_t^*$ , and  $i_t = i_t^* + \text{prem}_t$ .
  - $\Delta \bar{e} = 0$ : fixed exchange rate, inflation target consistent with the trend in the real exchange rate

# Modified Taylor Rule

- $c_1$  is the degree of control the central bank holds over the money market interest rate, higher value of  $c_1$  means a lower degree of control.
- Fixed exchange rate,  $\Delta \bar{e} = 0$ , since rational market agents know that the exchange rate will not change in the future, indicating a forward-looking behaviour.
- **Long-run average value of inflation represented by  $\bar{\pi}$  depends upon the foreign inflation, implying that  $\bar{\pi} = \bar{\pi}^* - \Delta \bar{q}$ .**
- $\Delta \bar{q}$  is the deviation in the REER equilibrium, such that  $q_t = e_t + p_t^* - p_t$ .
- $i_t^n$  is the neutral nominal interest rate, which is obtained as a trend real interest rate plus model-consistent inflation expectations
- The inflation target is consistent with a trend in the real exchange rate as represented in our model by the coefficient  $\bar{q}$ , limiting the central bank's choice either controlling the exchange rate or the interest rate.
- **$\bar{\pi}$  in equation is replaced with  $\bar{\pi}^* - \Delta \bar{q}$  during estimation.**



# Modified Taylor Rule - Estimation

- Modified Taylor rule as discussed before ((Berg et al. 2006a, 2006b) and Bulir (2014) is further simplified into a testable format:
  - Simulated risk premium data for the 'prem' variable adopted from Bhatta et al (2021a), monthly data averaged for quarterly.
  - Based on the findings of Bhatta et al (2021b), we assigned the negligible weight of domestic economic conditions and substantial weight of foreign, as follows:

$$i_t = 0.9(\alpha i_t^* + \beta prem_t) + 0.1(\rho i_{t-1} + (1 - \rho)(\gamma i_t^n + \delta \pi_{t+1}^e - \lambda \bar{\pi}^* + \Psi \Delta \bar{q} + \varphi \hat{y}_t)) + \varepsilon_t^i$$

## EXTENSION

- adding Nepal's stock exchange index and broad money supply, following the approach of Castro (2011) and Hsing (2009).

$$i_t = 0.9(\alpha i_t^* + \beta prem_t) + 0.1(\rho i_{t-1} + (1 - \rho)(\gamma i_t^n + \delta \pi_{t+1}^e - \lambda \bar{\pi}^* + \Psi \Delta \bar{q} + \varphi \hat{y}_t + \tau_1 M2 + \tau_2 Stock)) + \varepsilon_t^i$$

# Data

- Quarterly time series data from 2000Q1 to 2020Q4
- Interest (policy) rate: 91- days T-bill assumed to be NRB's policy rate, same rate of India's taken
  - Belief: 91 days' T-bill rate can represent the closet trend in the absence of CB's policy rate
- Exchange rates: nominal is NEER and real is REER (REER is not inflation adjusted, but exogenous); taken from Darvas (2012)
- CPI, real GDP
- Risk premium: calibrated data adopted (from our earlier study is from Bhatta et al (2021a))

# Modelling Technique : GMM

- Hansen (1982)'s Generalised method of moments(GMM) applied for estimation Addresses the endogeneity problem:
  - interest rate, the dependent variable, directly influences the risk premium and affects inflation and the exchange rate
  - feedback of the policy actions complex and interactive, also highlighted by Clarida et al. (1998).

$$Q(\alpha, \beta, \rho, \gamma, \delta, \lambda, \Psi, \varphi) = \left( \frac{1}{N} \sum_i z_i \varepsilon_i \right)' W \left( \frac{1}{N} \sum_i z_i \varepsilon_i \right)$$

- Where,  $W$  is a weighting matrix with  $W = \hat{S}^{-1}$  and  $\hat{S} = \frac{1}{N} \sum_i \hat{\varepsilon}_i^2 z_i z_i'$

- Eight instruments included: lagged values of the variables, following the Clarida et al (1998)
  - Model is overidentified

# Findings

# Results

Long-term Coefficients		Estimates	Short-run (Smoothened) Coefficients	Estimates
$i_t^*$	$\alpha$	0.223 (0.219)	(0.9) $\alpha$	0.2007
$prem_t$	$\beta$	0.132 (0.092)	(0.9) $\beta$	0.1188
$i_{t-1}$	$\rho$	0.732*** (0.081)	(0.1) $\rho$	0.073***
$i_t^n$	$\gamma$	0.781*** (0.332)	0.1(1- $\rho$ )* $\gamma$	0.021***
$\pi_{t+1}^e$	$\delta$	-0.666** (0.336)	0.1(1- $\rho$ )* $\delta$	-0.018**
$\bar{\pi}^*$	$\lambda$	0.279* (0.158)	0.1(1- $\rho$ )* $\lambda$	0.0075*
$\Delta \bar{q}$	$\Psi$	3.745** (1.689)	0.1(1- $\rho$ )* $\Psi$	0.1004**
$\hat{y}_t$	$\varphi$	0.059** (0.027)	0.1(1- $\rho$ )* $\varphi$	0.0016**
Wald $\chi^2$		1444.9 (p-value:0.00)		
Adj. R <sup>2</sup>		0.8408		
Hansen's J Statistics		7.785 (p-value:0.2543)		

# Results: Extended Model

Long-term Coefficients		Estimates	Short-run Coefficients	Estimates
$i_t^*$	$\alpha$	0.221 (0.213)	$(0.9)\alpha$	0.1989
$prem_t$	$\beta$	0.105 (0.094)	$(0.9)\beta$	0.0945
$i_{t-1}$	$\rho$	0.732*** (0.069)	$(0.1)\rho$	0.0732***
$i_t^n$	$\gamma$	0.799** (0.332)	$0.1(1-\rho)*\gamma$	0.021**
$\pi_{t+1}^e$	$\delta$	-0.652* (0.343)	$0.1(1-\rho)*\delta$	-0.017*
$\bar{\pi}^*$	$\lambda$	0.257 (0.161)	$0.1(1-\rho)*\lambda$	0.007
$\Delta\bar{q}$	$\Psi$	3.80** (1.83)	$0.1(1-\rho)*\Psi$	0.102**
$\hat{y}_t$	$\varphi$	0.059** (0.027)	$0.1(1-\rho)*\varphi$	0.002**
	$\tau_1$	-0.078* (0.043)	$0.1(1-\rho)*\tau_1$	-0.002*
	$\tau_2$	-0.0086 (0.0089)	$0.1(1-\rho)*\tau_2$	-0.0002
Wald $\chi^2$		1782.4 (p-value:0.00)		
Adj. R <sup>2</sup>		0.843		
Hansen's J Statistics		7.651 (p-value:0.2648)		

# Results...

- Extended model: results obtained are similar to the earlier estimation
  - shows robustness and consistency of earlier results
  - We do not reject the null hypothesis for Hansen J statistics, ensuring instrument validity.
- To further validate the model and robustness check: two stages least squares (2SLS), model estimated.
  - Similar coefficient estimates obtained
- Risk premium's positive impact to interest rate confirms the earlier findings by Bhatta et al (2021a) but insignificant
  - A higher risk premium (positive value of the premium) means Nepal's interest rate should be higher than India's and vice versa
  - Premium is a simulated (unobserved), it is a control variable for the proposed reaction function
- Interest rate substantially explained by foreign factors
  - No direct impact of foreign interest rate but REER, with a similar finding of Mohanty and Klau (2005) and Kendall and Ng (2013),

# Discussions

- REER appreciation may have a similar impact as of domestic inflation expectations in the monetary policy rule.
  - Nepal receives substantial remittance
  - Studies such as by Barajas et al. 2011; Hassan & Holmes 2013 showed remittances cause the REER to appreciate
  - Hien et al. (2019) confirm that the 'Dutch disease' effect appears only if the remittances are spent on consumption, which directly related Nepal
  - Remittances may have been accommodated via the REER
- The REER appreciation causes an increase in long-term rates but it would only be partially reflected due to the weak transmission channel
  - REER appreciation will further deteriorating the external trade competitiveness,
- The negative impact of inflation expectations against the Taylor (1993) principle
  - an outward-looking strategy must be considered while setting the interest rate to enhance the effectiveness of the monetary policy transmission.



# Policy implications

- the policy rate should follow the reaction function like the one proposed in this study
  - helps to communicate the signals of inflation expectations to the market and guides the central bank's policy rate path.
- Being the most substantiated variable affecting the interest rate, the REER appreciation may have a similar impact to that of domestic inflation expectations in the monetary policy rule.
  - A sharp currency appreciation would have two effects: a rise in the interest rate and a loss of international competitiveness.
- Inflation-targeting monetary policy would not be feasible for Nepal, as explained by Suzuki (2019)
  - The role of inflation expectation is however may be fulfilled by the REER

# Way Forward

- Nepal's benefits should arise from introducing a rule-based policy to enhance the policy effectiveness in the first phase.
  - A replacement of the discretionary policy with a rule prevents political influence, further strengthening institutional and operational independence and addressing the time inconsistency problem.
- When the future short-term interest rate can be best predicted using the REER, inflation targeting can be equivalent to exchange rate targeting (Svensson 1997).
  - An area to explore for Nepal would be an exchange rate targeting monetary policy.

**THANK YOU**

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